Available online at http://scik.org

J. Math. Comput. Sci. 11 (2021), No. 4, 4841-4862

https://doi.org/10.28919/jmcs/5913

ISSN: 1927-5307

A NOVEL APPROACH FOR HIGH-PERFORMANCE HEAT INDEX FORECASTING FOR THE HOTTEST REGION IN THAILAND

PIYATIDA YODPIBUL, THAMMARAT PANITYAKUL, NOODCHANATH KONGCHOUY*

Division of Computational Science, Prince of Songkla University, Hat Yai, Songkla 90110, Thailand

Copyright © 2021 the author(s). This is an open access article distributed under the Creative Commons Attribution License, which permits

unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Abstract. In Thailand, the hottest area is the northern inland plains located in the Northern Thailand. In summer,

from the middle of February to the middle of May, the temperature may rises up to 40° or above due to the

changing of northeast monsoon to southwest monsoon and the impact of relative humidity. Relative humidity is

the major factor that makes people feel hotter than the actual temperature. If only the air temperature was noticed,

it is possible to take risk of overheating and heat illness, especially heat stroke that can be deadly. The heat index

has used as an effective warning measurement, this calculated by Steadman's equation and yields the real feel of

body. In order to prevent the heat illness, the predictive analytics such as time series forecasting should be applied.

The regular series was constructed by several time points in consecutive daily heat index, the seasonal and cycle

effects will be analyzed simultaneously. This scenario leads to the complicated time series model and may cause

inaccuracy of forecasting. The proposed study modify the data structure as the series of specific date and time for

thirty years, i.e., 1-April to 30-April at time 4.00 p.m., this reveals distinguished increasing trend from year by year.

Three trend-focused forecasting model be applied, the two benchmarking models are Holt's linear trend model and

time series regression model, being compared with the proposed model called autocorelated-based decomposition.

The forecasting results of Uttaradit and Chiang Mai provinces heat index in recent thirty years show that the

proposed approach yields more accuracy than the benchmarks. For Uttaradit, the MAPE of the proposed model

less than the others from 26.8% to 36.9%, and less than the others from 48.9% to 55.9% in RMSE. For Chiang

*Corresponding author

E-mail address: nootchanath.k@psu.ac.th

Received April 23, 2021

4841

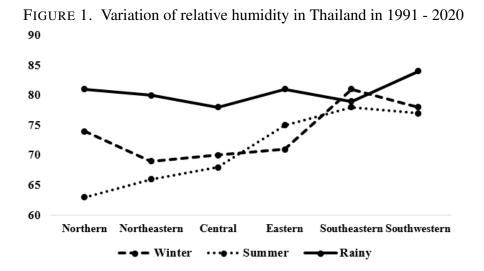
Mai, the MAPE of the proposed model less than the others from 16.9% to 36.9%, and less than the others from 27.5% to 61.6% in RMSE.

Keywords: Northern Thailand; heat index; Holt's linear trend model; time series regression; autocorrelation function; decomposition.

2010 AMS Subject Classification: 62P12.

1. Introduction

The hottest area in Thailand is the area that away from the coast called the northern inland plains, specifically, the Northern Thailand. In December and January the Northern Thailand has lowest temperature especially the night temperature that can drop to 5°C or below due to the high atmospheric pressure from China. Oppositely, in the summer period from February to May, the temperature may rises up to 40°C or above due to the changing of northeast monsoon to southwest monsoon The major factor that causes this region has highest fluctuations in yearly temperatures is the impact of relative humidity. Figure 1 shows the relative humidity in Thailand for last thirty years (1991-2020), we noticed that the northern inland region has highly variation compared to other regions. The Southeastern and Southwestern have quite stable relative hu-



midity that leads to the small changing in temperature. While the Northern has wider spread of relative humidity than the others, this leads the summer temperature of this region higher than others, especially in Uttaradit province which has highest temperature in Thailand. In addition,

the relative humidity is the major factor that makes people feel hotter than the actual temperature. Because the increasing relative humidity, the decreasing rate of sweating, so it actually feels warmer outside than it is. If only the air temperature was noticed, it is possible to take risk of overheating and heat illness, especially heat stroke that can be deadly [20]. In order to prevent the illness from hot weather, we suppose to know the apparent temperature or real feeling temperature instead of the air temperature. The type of temperature called Heat Index, i.e., the combination between air temperature and relative humidity, used for measure the actual feel temperature. Hot weather leads to illnesses, for instance, heat syncope, heat cramps, heat exhaustion and heat stroke [2], [3], [19] not only for the the tourists but for the workers also as mention in [4] and [18]. Furthermore, the information be used as a scientific tool for helping a heat warning system construction [7] and [8], that will be used as a helping planning tool of heat in vulnerable areas as mention in [4] and [5]. This study provides a high performance heat index forecasting technique called autocorrelation-based decomposition (ACD) comparing to the Holt's linear trend model (HLT) [10], [11] and classical decomposition model (CDM) [16], [17].

2. DATA AND METHOD

Uttaradit province (UT) and Chiang Mai province (CM) are the northern inland region of thailand. From the period 2001 to 2016, the maximum temperatures in Thailand rose from 38–41° to 42–44° as mention in [4] and [7]. This suitable for heat index studying, because Uttaradit province was the hottest province in Thailand in many recent years, and Chiang Mai was the famous city for the tourists from the temperate zone countries and has the international airport. Forecasting heat index in both provinces has highly advantage for preventing the illness from hot weather. Since the heat index is the combination between air temperature and relative humidity, the Steadman's equation [9] and [22] describes the heat index calculation as shown in (1)

$$HI = 42.38 + 2.049T + 10.14RH + 0.2248TRH + 6.88 * 10^{-3}T^{2} +$$

$$= 5.482 * 10^{-2}RH^{2} + 1.228 * 10^{-3}T^{2}RH + 8.528 * 10^{-4}TRH^{2} +$$

$$= 1.99 * 10^{-6}T^{2}RH^{2}$$

where HI: heat index (${}^{\circ}F$); T: dry bulb temperature (${}^{\circ}F$); RH: relative humidity (%).

In Thailand, there was only air temperature data from the weather stations or meteorological station, we have to apply the Steadman's equation [9] in order to compute the heat index of Uttaradit and Chiang Mai provinces, and combine the heat index data as the original series. There are 30-years time series of daily heat index between April 15th and May 15th, the hottest period in Thailand, was collected. It is reasonable for consider this period because there is the highest temperature period in every regions of Thailand, especially in the afternoon, says, 4.00 pm. Another reason is avoiding the complicated time series models for predicting the consecutive data points in this period and being comprised year after year. Becaused of these reasons, we have 30 series of heat index to forecast, i.e., series of heat index on April 15th at time 4.00 p.m from 1990 to 2020, series of heat index on April 16th at time 4.00 p.m. from 1991 to 2020, ..., and so on, as shown in Table 1.

TABLE 1. 1991-2020 series of heat index on the same date and time from 1-April through 30-April at time 4.00 p.m. of Uttaradit and Chiang Mai provinces

Year		adit		Chiang Mai				
	1-Apr	2-Apr		30-Apr	1-Apr	2-Apr		30-Apr
1991	32.2	32.7		41.1	31.6	29.5		34.6
1992	41.4	43.9		46.3	37.2	37.2		31.4
1993	42.3	41.9		42.8	36.0	34.8		35.6
1994	40.1	43.5		47.2	32.7	29.8		36.1
1995	42.2	45.4		45.5	36.3	36.2		36.6
÷	÷	÷		÷	÷	÷		÷
2018	39.2	38.1		38.2	32.1	34.7		34.0
2019	40.5	40.8		43.4	36.4	36.5		40.6
2020	43.2	42.7		43.2	38.2	37.6		32.8

Table 2-3 represent the essential statistics which explain natural characteristics of the thirty series from Uttaradit and Chiang Mai provinces. The mean values of all series from any

province are most likely the same as well as standard deviation values, this states that the most of series are from the similar distribution.

While Figure 2-3 show the physical of essential statistics of Uttaradit and Chiang - Mai heat index in April day by day.

FIGURE 2. Minimum, Mean and Maximum values in each day of April of Uttaradit heat index in 1991 - 2020

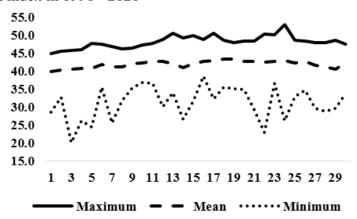


FIGURE 3. Minimum, Mean and Maximum values in each day of April of Chiang Mai heat index in 1991 - 2020

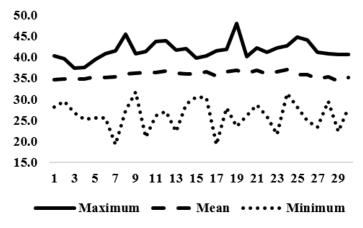


Figure 4-5 show the box-plots of Uttaradit and Chiang Mai heat index of each day in April.

Figure 2-5 show the distribution of each day heat index in April, we noticed that the central tendency of 30 days are quite similar, so, if we use the average of each series represent the whole series we will misunderstand the nature of the series. As well as the minimum of each series,

TABLE 2. Essential statistics of 1991-2020 heat index series in each day of April from Uttaradit provinces

Date	Min	Max	Range	Mean	Median	S.D.
1-Apr	28.47	44.91	16.44	39.85	40.79	3.63
2-Apr	32.71	45.41	12.70	40.32	41.55	3.10
3-Apr	20.07	45.64	25.57	40.60	41.15	4.54
4-Apr	26.22	45.92	19.70	40.67	41.01	3.65
5-Apr	24.38	47.61	23.23	40.71	41.55	4.13
6-Apr	35.57	47.36	11.79	41.74	42.26	2.87
7-Apr	25.45	46.89	21.44	41.15	41.71	4.01
8-Apr	31.91	46.13	14.22	41.21	41.49	3.33
9-Apr	35.18	46.37	11.19	42.06	42.60	2.96
10-Apr	36.88	47.19	10.31	42.32	42.51	2.86
11-Apr	36.52	47.73	11.21	42.72	43.50	3.12
12-Apr	30.06	48.65	18.59	42.63	42.82	3.92
13-Apr	34.04	50.46	16.42	42.07	42.25	4.03
14-Apr	26.56	49.12	22.56	40.89	41.88	5.26
15-Apr	31.62	49.83	18.21	42.06	42.62	3.90
16-Apr	38.47	48.77	10.30	42.64	42.08	2.60
17-Apr	32.26	50.52	18.26	42.94	42.81	3.39
18-Apr	35.49	48.52	13.03	43.22	43.19	3.10
19-Apr	35.16	47.87	12.70	43.34	43.38	2.69
20-Apr	34.85	48.27	13.42	42.63	42.73	2.85
21-Apr	29.26	48.36	19.10	42.68	43.38	4.39
22-Apr	22.82	50.20	27.38	42.47	43.32	5.15
23-Apr	36.52	50.02	13.51	42.71	42.85	3.44
24-Apr	26.21	52.91	26.70	42.97	43.42	4.80
25-Apr	32.82	48.55	15.73	42.24	43.01	3.86
26-Apr	34.59	48.36	13.77	42.77	43.16	3.54
27-Apr	29.55	47.84	18.29	41.51	42.43	4.59
28-Apr	28.67	47.84	19.17	41.17	41.86	4.33
29-Apr	29.74	48.61	18.87	40.51	41.28	4.96
30-Apr	33.66	47.44	13.78	42.14	42.66	3.43

TABLE 3. Essential statistics of 1991-2020 heat index series in each day of April from Uttaradit provinces

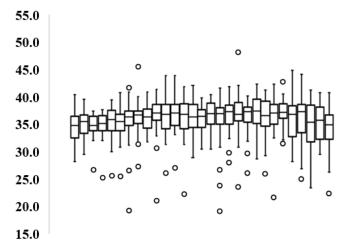
Date	Min	Max	Range	Mean	Median	S.D.
1-Apr	28.21	40.39	12.18	34.61	34.76	2.70
2-Apr	29.46	39.60	10.15	34.89	35.46	2.52
3-Apr	26.69	37.45	10.76	34.81	34.81	2.14
4-Apr	25.27	37.64	12.37	34.87	35.15	2.34
5-Apr	25.57	39.45	13.88	35.29	35.85	2.89
6-Apr	25.50	40.81	15.31	35.28	35.53	2.85
7-Apr	19.22	41.65	22.43	35.35	36.29	4.19
8-Apr	27.32	45.46	18.15	36.04	36.66	3.44
9-Apr	31.77	40.89	9.12	36.21	36.30	2.37
10-Apr	20.97	41.37	20.40	36.35	37.09	3.88
11-Apr	26.10	43.86	17.76	36.44	36.74	3.45
12-Apr	27.05	43.90	16.84	36.71	37.07	3.08
13-Apr	22.26	41.78	19.52	36.15	36.79	3.78
14-Apr	28.87	42.07	13.19	36.12	36.34	3.00
15-Apr	30.49	39.81	9.32	36.03	36.50	2.51
16-Apr	30.42	40.43	10.02	36.53	36.94	2.57
17-Apr	19.12	41.58	22.46	35.53	36.88	4.90
18-Apr	28.03	41.89	13.86	36.66	37.24	3.09
19-Apr	23.50	48.11	24.61	36.86	36.81	4.02
20-Apr	26.10	40.18	14.08	36.29	37.30	3.00
21-Apr	28.60	42.32	13.71	36.96	37.43	3.04
22-Apr	25.89	41.27	15.38	36.02	36.50	3.82
23-Apr	21.59	42.30	20.71	36.61	37.09	3.70
24-Apr	31.43	42.69	11.26	37.05	37.17	2.56
25-Apr	28.11	44.83	16.73	35.86	36.78	3.90
26-Apr	24.93	44.11	19.17	35.84	37.33	4.24
27-Apr	23.33	41.23	17.90	34.85	35.30	4.20
28-Apr	29.45	40.81	11.36	35.29	35.78	3.18
29-Apr	22.34	40.78	18.44	34.25	34.92	4.14
30-Apr	28.01	40.64	12.63	35.26	35.12	3.02

55.0 50.0 45.0 40.0 35.0 30.0 25.0

15.0

FIGURE 4. Box plots of all heat index series in April 1991 - 2020 of Uttaradit

FIGURE 5. Box plots of all heat index series in April 1991 - 2020 of Chiang Mai



the high variation of these values may causes the wrong explanation of the whole series but we do not interested in minimum values. For the maximum values of each days, we observed that Uttaradit has lower variation than Chiang Mai, this may leads us to the considerably different final models between the two provinces.

As the reason of avoiding the complicated time series models, these models including seasonality that is the major task to handle. By using time series structure as we mention above, the series be remained only the trend, cycle and irregular components. It is appropriate to apply time series analysis methods which focus on trend.

We applied HLT and CDM as benchmark models, the HLT can be considered separately as follows,

$$\hat{z}_{i+l|i} = a_i + lm_i$$

(3)
$$a_i = \vartheta z_i + (1 - \vartheta)(a_{i-1} + m_{i-1})$$

(4)
$$m_i = \zeta(a_i - a_{i-1}) + (1 - \zeta)m_{i-1}$$

(7)

where a_i stands for the level estimate at time i, m_i is a level estimate at time i, ϑ is smoothing parameter of the level, ζ is the smoothing parameter of the trend which $0 \le \vartheta \le 1$ and $0 \le \zeta \le 1$.

We applied additive classical decomposition method (CDM) as another benchmarking model in oder to separate the time series into linear trend and seasonal components, as well as error, and to provide forecasts. The additive model be applied for both CDM and ACD because the magnitude of the seasonal pattern does not change as the series goes up or down. Any time series has four components, i.e., the long-term tendency called trend, the periodic fluctuation within a certain time period called seasonality, the periodic fluctuation over a large time interval called cycles and the random noise or error called irregular.

For both CDM and ACD started with trend estimation, \hat{T}_i , that calculated by s-moving average since s stands for seasonal periods. (As we mentioned above, the data structure of this study designed for avoiding seasonality, seasonality in those data set might be psuedo-seasonality that explained by average El Nino cycle year in eastern tropical Pacific, that is five years.) The second step is detrending by subtraction the original series with trend estimates, obtaining the detrended series, $z_i - \hat{T}_i$. The third step is computing the seasonal component by averaging the detrended values of each season, obtaining the seasonal indices, \hat{S}_i . The fourth step is subtracting the original series by the seasonal indices, obtaining the seasonal adjusted data, W_i . The fifth step is predicting W_i by simple prediction method such as random walk with drift (applied for CDM), obtaining the \hat{W}_i . The final step is just adding \hat{W}_i with \hat{S}_i to obtain the predicted \hat{z}_i . The fifth step has changed for ACD by adoption the autoregressive moving average (ARMA) approach, we observed that the ARMA(0,1) or MA(1) is the best condition for \hat{T}_i , this is an ensemble probabilistic approach [21]. The ARMA(p, q) in backshift form shown in (5) -

$$\phi_p(B)z_i = \theta_q(B)\varepsilon_i$$

$$\phi_p(B) = 1 - \phi_1 B - \ldots - \phi_p B^p$$

(7)
$$\theta_a(B) = 1 - \theta_1 B - \ldots - \theta_a B^q$$

where ϕ and p represent the parameter and the order of autoregressive process respectively, θ and q represent the parameter and the order of moving average process respectively, and B is backshift operator which $B^k(z_i) = z_{i-k}$.

This study had proposed the new approach of time series decomposition method called autocorrelation-based decomposition, ACD, since the autocorrelation means self-linear relationship between the original series and its lags, as shown in (8), the sample autocorrelation of any time series,

(8)
$$r_k = \frac{\sum_{i=k+1}^n (z_i - \bar{z})(z_{i-k} - \bar{z})}{\sum_{i=1}^n (z_i - \bar{z})^2},$$

where n is the length of time series (in this study all series have length 30). If the magnitude of r_k not over the lower or upper bound of correlogram, then we can conclude that no linear relationship between the original series and k^{th} -lag. Our ascription is at the same time (e.g. Apr 18th, 4.00 p.m.) the heat index increasing linearly year after year. The classical decomposition consider a time series by components, eventhough the seasonality was ignored by the data collecting structure, but the seasonality adjustment will screen the parsimonious seasonality from the original series. The cycle component be composed of the remainder component as well as the irregular part. We modified the autocorrelation function as the predictive tool for the remainder component, the component after trend and seasonality decomposed, of the classical decomposition method.

The mean absolute percentage error (MAPE) and root mean square error be used as the performance meansurement of all models, where MAPE (%) measures the accuracy of fitted time series values, and RMSE measures the accuracy of fitted time series values. These two

measurements computed from the testing set of the data (2016 -2020) while the training set is according data series in 1991 -2015

The model validation has tested by Ljung-Box test, which tests the staionarity of the residuals by applying the Ljung-Box Q statistics in (9) that comprised from the autocorrelation function of the residuals.

(9)
$$LBQ = n(n+2) \sum_{j=1}^{k} (n-j)^{-1} r_k^2$$

where r_k^2 is the autocorrelation at k^{th} -lag.

3. RESULTS AND DISCUSSION

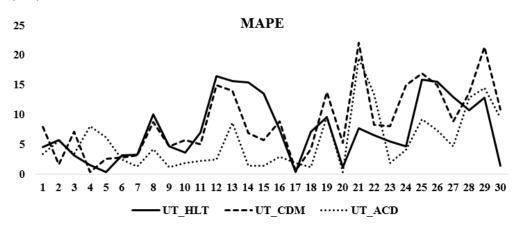
All models performance are shown in Figure 6 - 7, For Uttaradit, the performance of ABD explicitly better than HLT and CDM in both MAPE and RMSE, while it is not differ from CDM in both MAPE and RMSE in Chiang Mai. In an average manner, as shown in Table 4, ADC manifestly overcome HLT and CMD by 27% - 37% reduction in MAPE and 49% - 56% reduction in RMSE for Uttaradit. As well as for Chiang Mai, the percentage reduction is 17% - 37% in MAPE and 28% - 62% in RMSE.

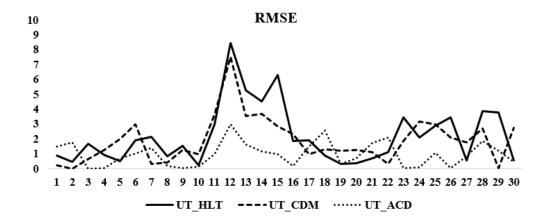
TABLE 4. Heat index forecasting performance of Uttaradit and Chiang Mai provinces

		Ut	taradit	Chiang Mai		
		Average	%Reduction	Average	%Reduction	
MAPE(%)	HLT	7.57 26.76		7.74	36.99	
	CDM	8.79 36.91		5.86	16.86	
	ACD	5.55		4.87		
RMSE	HLT	2.21	55.99	2.45	61.62	
	CDM	1.91 48.99		1.30	27.50	
	ACD	0.97		0.94		

Another visible evidence are from the residual diagnostic as demonstrated in Figure 8. Since the model validation testing by comparing the Ljung-Box Q statistics in (9) to the χ^2 -critical value which 1-degree of freedom, none of ACD over the critical value in both two provinces, we can clearly conclude that ACD suitable to forecast all series and area.

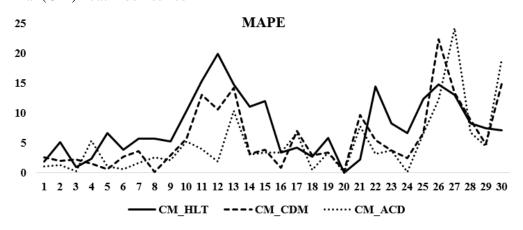
FIGURE 6. MAPE and RMSE values from all forecasting methods for Uttaradit (UT) heat index series





In order to perform forecasting and comparison between all models, only the series that valid in all methods were selected, there are eighteen series remaining, i.e., 1-Apr, 3-Apr, 5-Apr, 6-Apr, 8-Apr, 11-Apr, 12-Apr, 13-Apr, 15-Apr, 16-Apr, 17-Apr, 18-Apr, 20-Apr, 22-Apr, 23-Apr, 26-Apr, 27-Apr, 29-Apr and 30-Apr. The eightteen series be forecasted in 3-steps ahead and forecasting interval be computed as well, as shown in Table 5 - 6. Figure 9-10 show that the ACD forecasts and their upper bound totally less than HLT and CDM methods, while the lower bound values in Figure 11 quite similar for all methods. This support the superior of ACD method. Afterall, the 3-steps ahead forecasting of series 27-Apr (the hottest date for recent many years) with ACD and their intervals be compared to the other intervals, as shown in Figure 12-13. The forecasting interval of ACD narrower and smoother than HLT and CDM, this seems to be a good characteristic of any forecasting models.

FIGURE 7. MAPE and RMSE values from all forecasting methods for Chiang Mai (CM) heat index series



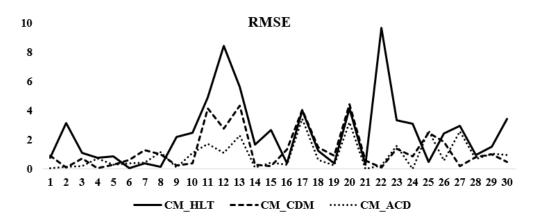
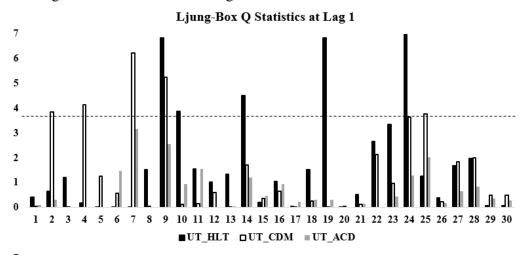


FIGURE 8. Ljung-Box Q statistics for model validation testing of Uttaradit and Chiang Mai heat index forecasting models



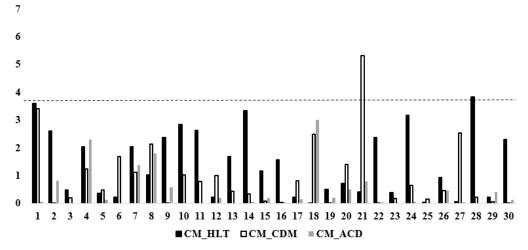


TABLE 5. Three steps ahead forecasts of the series on last week of April in Uttaradit

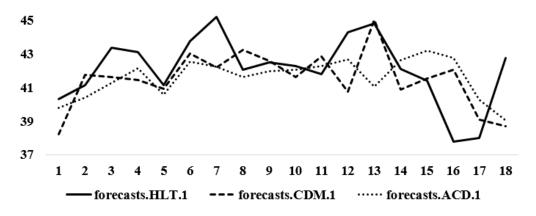
-									
HLT	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	44.82	44.99	45.16	54.39	55.02	55.69	35.26	34.97	34.63
23-Apr	42.10	42.09	42.09	50.15	50.76	51.45	34.06	33.42	32.72
26-Apr	41.38	41.29	41.19	49.59	50.11	50.69	33.18	32.46	31.68
27-Apr	37.78	37.41	37.03	47.10	47.36	47.70	28.46	27.45	26.37
29-Apr	37.98	37.73	37.48	48.33	49.19	50.17	27.63	26.28	24.80
30-Apr	42.76	42.80	42.85	50.15	51.77	53.54	35.37	33.84	32.16
CDM	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	45.05	46.06	43.76	54.75	59.77	60.55	35.35	32.34	26.96
23-Apr	40.86	42.85	40.82	47.30	51.95	51.96	34.43	33.75	29.67
26-Apr	41.55	39.90	42.15	48.24	49.35	53.72	34.87	30.45	30.57
27-Apr	42.06	42.14	41.57	50.79	54.48	56.68	33.33	29.80	26.46
29-Apr	39.10	38.95	37.34	48.88	52.78	54.28	29.32	25.11	20.40
30-Apr	38.68	40.66	38.09	44.79	49.30	48.68	32.56	32.01	27.51
ACD	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	41.03	40.03	43.12	50.89	53.98	60.21	31.16	26.08	26.03
23-Apr	42.63	40.23	43.32	49.16	49.48	54.64	36.09	30.99	32.01
26-Apr	43.19	40.27	43.36	49.77	49.57	54.75	36.61	30.96	31.96
27-Apr	42.74	39.08	42.17	50.61	50.23	55.82	34.86	27.94	28.53
29-Apr	40.27	37.64	40.73	49.92	51.29	57.44	30.62	23.99	24.02
30-Apr	39.02	39.46	42.55	43.69	46.07	50.64	34.35	32.86	34.46

TABLE 6. Three steps ahead forecasts of the series on last week of April in Chiang Mai

HLT	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	36.27	36.31	36.35	44.30	44.72	45.17	28.25	27.91	27.54
23-Apr	39.23	39.46	39.69	46.10	46.97	47.91	32.37	31.95	31.47
26-Apr	33.70	33.54	33.38	42.50	43.26	44.12	24.91	23.82	22.64
27-Apr	33.18	32.95	32.72	43.06	44.82	46.77	23.30	21.08	18.68
29-Apr	31.82	31.58	31.33	40.63	41.25	41.97	23.01	21.90	20.70
30-Apr	34.72	34.71	34.70	41.87	42.35	42.88	27.57	27.06	26.51
CDM	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	36.20	38.16	35.77	43.36	48.28	48.17	29.04	28.04	23.38
23-Apr	38.77	37.36	36.98	45.89	47.43	49.31	31.64	27.29	24.64
26-Apr	31.80	33.74	35.02	39.73	44.95	48.75	23.87	22.53	21.30
27-Apr	36.93	38.20	37.12	44.78	49.30	50.71	29.08	27.10	23.53
29-Apr	30.38	31.50	31.54	38.06	42.36	44.84	22.70	20.64	18.24
30-Apr	33.97	32.12	33.46	39.64	40.14	43.28	28.30	24.10	23.65
ACD	forecast 1	forecast 2	forecast 3	upper 1	upper 2	upper 3	lower 1	lower 2	lower 3
22-Apr	35.98	33.63	36.72	43.21	43.86	49.25	28.75	23.41	24.20
23-Apr	36.33	34.15	37.24	43.43	44.20	49.55	29.22	24.10	24.93
26-Apr	35.56	33.22	36.31	43.87	44.97	50.70	27.25	21.47	21.92
27-Apr	30.18	32.12	35.21	37.40	42.33	47.71	22.97	21.92	22.71
29-Apr	35.25	31.67	34.76	42.66	42.15	47.59	27.84	21.19	21.92
30-Apr	35.21	32.87	35.96	41.06	41.14	46.09	29.36	24.60	25.83

 $FIGURE\ 9.\ Comparisons\ of\ one-step\ ahead\ forecasts\ from\ all\ models$

Uttaradit



Chiang Mai

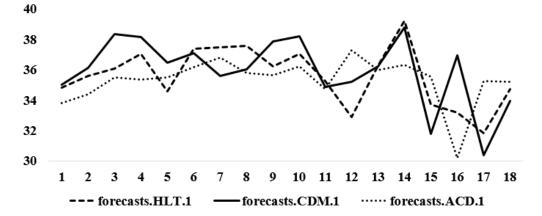
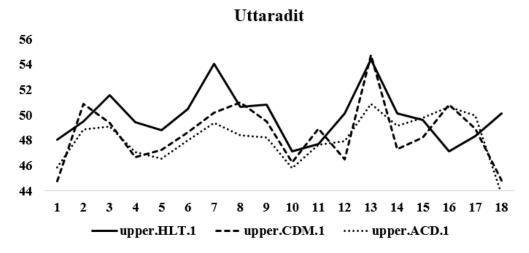


FIGURE 10. Comparisons of upper bound of one-step ahead forecasts from all models



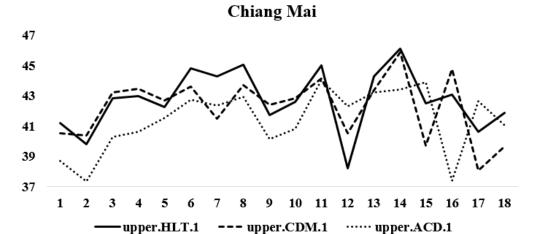


FIGURE 11. Comparisons of lower bound of one-step ahead forecasts from all models

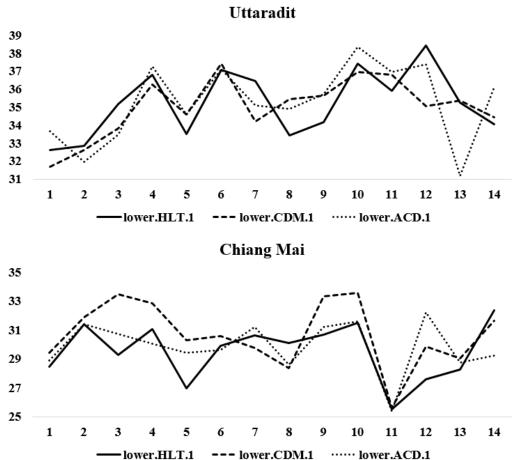
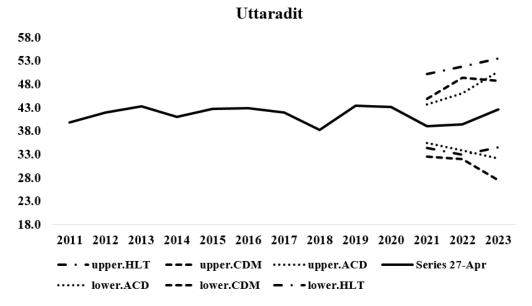


FIGURE 12. Three-steps ahead forecasts of series 27-Apr in Uttaradit



Chiang Mai

48

43

38

33

28

2011 2012 2013 2014 2015 2016 2017 2018 2019 2020 2021 2022 2023

- · - upper.HLT --- upper.CDM ······ upper.ACD — Series 27-Apr

····· lower.ACD --- lower.CDM - · - lower.HLT

FIGURE 13. Three-steps ahead forecasts of series 27-Apr in Chiang Mai

The proposed approach for heat index forecasting including data gathering, series construction and the modified decomposition method, has delivered the better forecasting result comparing to the classical benchmarking models, HLT and CDM, not only by performance measurements and model validity, but future forecasting also. However, the further study should be invent in other approach for instance, frequency domain approach or machine learning approach.

CONFLICT OF INTERESTS

The author(s) declare that there is no conflict of interests.

REFERENCES

- [1] R. Kotharkar, A. Ghosh, V. Kotharkar, Estimating summertime heat stress in a tropical Indian city using Local Climate Zone (LCZ) framework, Urban Climate, 36 (2021), 100784.
- [2] E.A. Lehnert, G. Wilt, B. Flanagan, E. Hallisey, Spatial exploration of the CDC's Social Vulnerability Index and heat-related health outcomes in Georgia, Int. J. Disaster Risk Reduction. 46 (2020), 101517.
- [3] F. Golbabaei, H. Heidari, A. Shamsipour, A.R. Forushani, A. Gaeini, A new outdoor environmental heat index (OEHI) as a simple and applicable heat stress index for evaluation of outdoor workers, Urban Climate, 29 (2019), 100479.

- [4] W. Phanprasit, K. Rittaprom, S. Dokkem, A.C. Meeyai, V. Boonyayothin, J.J.K. Jaakkola, S. Näyhä, Climate Warming and Occupational Heat and Hot Environment Standards in Thailand, Safety Health Work, 12 (2021), 119-126.
- [5] S.G. Nayak, S. Shrestha, P.L. Kinney, et al. Development of a heat vulnerability index for New York State, Public Health, 161 (2018), 127-137.
- [6] J. Song, B. Huang, J.S. Kim, J. Wen, R. Li, Fine-scale mapping of an evidence-based heat health risk index for high-density cities: Hong Kong as a case study, Sci. Total Environ. 718 (2020), 137226.
- [7] S.O. Stapleton, L. Sabbag, K. Hawley, P. Tran, L. Hoang, P.H. Nguyen, Heat index trends and climate change implications for occupational heat exposure in Da Nang, Vietnam, Climate Services, 2–3 (2016), 41-51.
- [8] Y.T. Cheng, S.C.C. Lung, J.S. Hwang, New approach to identifying proper thresholds for a heat warning system using health risk increments, Environ. Res. 170 (2019), 282-292.
- [9] R.G. Steadman, The assessment of sultriness. Part I: A temperature-humidity index based on human physiology and clothing science, J. Appl. Meteorol. Climatol. 18 (1979), 861-873.
- [10] L. Liu, L. Wu, Predicting housing prices in China based on modified Holt's exponential smoothing incorporating whale optimization algorithm, Socio-Econ. Plan. Sci. 72 (2020), 100916.
- [11] A.L.S. Maia, F.A.T. Carvalho, Holt's exponential smoothing and neural network models for forecasting interval-valued time series, Int. J. Forecast. 27 (2011), 740-759.
- [12] W. Chen, H. Xu, Z. Chen, M. Jiang, A novel method for time series prediction based on error decomposition and nonlinear combination of forecasters, Neurocomputing, 426 (2021), 85-103.
- [13] C.E. Montagnon, Forecasting by splitting a time series using Singular Value Decomposition then using both ARMA and a Fokker Planck equation, Physica A: Stat. Mech. Appl. 567 (2021), 125708.
- [14] M. Zune, L. Rodrigues, M. Gillott, The vulnerability of homes to overheating in Myanmar today and in the future: A heat index analysis of measured and simulated data, Energy Buildings, 223 (2020), 110201.
- [15] G.F. Ghalhari, S.F. Dehghan, F. Shakeri, M.J. Ghanadzadeh, M. Asghari, Assessing the monthly changes of heat stress indices in outdoor environment during a 15-year period: Case of study in a dry and warm climate, Urban Climate, 31 (2020), 100538.
- [16] W. Chen, H. Xu, Z. Chen, M. Jiang, A novel method for time series prediction based on error decomposition and nonlinear combination of forecasters, Neurocomputing, 426 (2021), 85-103.
- [17] W. Chen, H. Xu, Z. Chen, M. Jiang, C.E. Montagnon, Forecasting by splitting a time series using Singular Value Decomposition then using both ARMA and a Fokker Planck equation, Physica A: Stat. Mech. Appl. 567 (2021), 125708.
- [18] S. Heo, M.L. Bell, J.T. Lee, Comparison of health risks by heat wave definition: Applicability of wet-bulb globe temperature for heat wave criteria, Environ. Res. 168 (2019), 158-170.

- [19] G.F. Ghalhari, S.F. Dehghan, F. Shakeri, M.J. Ghanadzadeh, M. Asghari, Assessing the monthly changes of heat stress indices in outdoor environment during a 15-year period: Case of study in a dry and warm climate, Urban Climate, 31 (2020), 100538.
- [20] T.I. Sung, P.C. Wu, S.C. Lung, C.Y. Lin, M.J. Chen, H.J. Su, Relationship between heat index and mortality of 6 major cities in Taiwan, Sci. Total Environ. 442 (2013), 275-281.
- [21] A. Dube, H. Singh, R. Ashrit, Heat Waves in India during MAM 2019: Verification of ensemble based probabilistic forecasts and impact of bias correction, Atmospheric Res. 251 (2021), 105421.
- [22] S. Heo, M. L. Bell, J. T. Lee, Comparison of health risks by heat wave definition: Applicability of wet-bulb globe temperature for heat wave criteria, Environ. Res. 168 (2019), 158-170.